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Author manuscript

*J Am Acad Child Adolesc Psychiatry*. Author manuscript; available in PMC 2017 February 01.

Published in final edited form as:

*J Am Acad Child Adolesc Psychiatry*. 2016 February ; 55(2): 137–145. doi:10.1016/j.jaac.2015.11.011.

## Mode of Anisotropy Reveals Global Diffusion Alterations in Attention-Deficit/Hyperactivity Disorder

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### Abstract

**Objective**—Diffusion tensor imaging (DTI) can identify structural connectivity alterations in attention-deficit/hyperactivity disorder (ADHD). Most ADHD DTI studies have concentrated on regional differences in fractional anisotropy (FA) despite its limited sensitivity to complex white matter architecture and increasing evidence of global brain differences in ADHD. Here, we examine multiple DTI metrics in separate samples of children and adults with and without ADHD with a principal focus on global between-group differences.

**Method**—Two samples: adults with ADHD ( $n = 42$ ) and without ( $n = 65$ ) and children with ADHD ( $n = 82$ ) and without ( $n = 80$ ) were separately group matched for age, sex, and head motion. Five DTI metrics (FA, axial diffusivity, radial diffusivity, mean diffusivity, and mode of anisotropy) were analyzed via tract-based spatial statistics. Group analyses tested for diagnostic differences at the global (averaged across the entire white matter skeleton) and regional level for each metric.

**Results**—Robust global group differences in diffusion indices were found in adults, with the largest effect size for mode of anisotropy (MA; Cohen's  $d = 1.45$ ). Global MA also differed significantly between groups in the pediatric sample ( $d = 0.68$ ). In both samples, global MA increased classification accuracy compared to the model with clinical Conners' ADHD ratings alone. Regional diagnostic differences did not survive familywise correction for multiple comparisons.

**Conclusion**—Global DTI metrics, particularly the mode of anisotropy, which is sensitive to crossing fibers, capture connectivity abnormalities in ADHD across both pediatric and adult

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Dr. Reiss served as the statistical expert for this research.

Disclosure: Drs. Yoncheva, Reiss, Kelly, Di Martino, Lazar, Zhou, Milham, Castellanos, and Mr. Somandepalli report no biomedical financial interests or potential conflicts of interest.

samples. These findings highlight potential diffuse white matter microarchitecture differences in ADHD.

### Keywords

ADHD; DTI; fractional anisotropy; adults; biomarkers

Models of the pathophysiology of attention-deficit/hyperactivity disorder (ADHD), a childhood-onset psychiatric disorder, have evolved from concentrating on fronto-striatal-cerebellar circuits to encompassing large-scale distributed networks.<sup>1-3</sup> Diffusion tensor imaging (DTI), which allows quantification of white matter microstructure, can inform the whole-brain substrates of pathologic alterations in structural connectivity. Most DTI studies of ADHD have limited their scope to tracts selected a priori or are pending definitive replication with rigorous control for multiple comparisons<sup>4</sup>; consequently, the localization and significance of particular regional white matter abnormalities in ADHD remain tentative.

In considering the heterogeneity of DTI findings, we note the global structural alterations in ADHD that other imaging modalities have consistently detected. Reliable overall reductions in total brain volume<sup>5-7</sup> and global cortical thickness<sup>8-10</sup> led us to reason that white matter connectivity might also be globally altered in ADHD. Furthermore, although delineating regional DTI differences can isolate particular loci of ADHD-related abnormalities,<sup>11-14</sup> global measures reduce neuroimaging data dimensionality, which is statistically advantageous,<sup>15</sup> providing impetus for a systematic examination of global diffusion indices.

The most ubiquitously reported diffusion parameter is fractional anisotropy (FA).<sup>16</sup> Although FA is often interpreted as indexing white matter “integrity,” Jones *et al.*<sup>17</sup> compellingly argued that this is an oversimplification. Moreover, interpreting FA is particularly problematic in areas with complex white matter architecture, e.g., crossing fibers, the proportion of which ranges from 63% to 90% in typical-resolution white matter voxels.<sup>18</sup> As the precise nature of white matter pathology in ADHD is unclear, complementary indices of diffusion tensor geometry warrant investigation.

One candidate tensor shape metric is the mode of anisotropy (MA),<sup>19</sup> not to be confused with the statistical term mode denoting the most frequent item in a set. MA is mathematically orthogonal to FA and quantifies second-order geometric properties,<sup>19,20</sup> notably resolving whether anisotropy is more planar (e.g., due to predominant crossing fibers within a voxel) or more linear (see Figure S1, available online). Investigators have begun to examine MA in brain disorders.<sup>21-23</sup> MA has contributed unique information relevant to clinical neurodegenerative progression,<sup>24</sup> but it has yet to be examined in ADHD.

The current study investigated global white matter microstructure in individuals with ADHD, assessing a gamut of DTI indices (including MA and conventional FA). Primary analyses were conducted in an adult sample and, to assess generalizability, repeated in a separate, pediatric sample with data acquired using the same scanner and imaging protocol. Within each of the 2 samples, the ADHD group was contrasted with an age-matched

neurotypical (NT) comparison group. Tract-based spatial statistics (TBSS)<sup>25</sup> was used to create a white matter skeleton common to all individuals in each sample. First, all global measures, aggregated across all white matter voxels in the TBSS skeleton, were contrasted between ADHD and NT groups. Second, we explored separately in each sample the prediction of ADHD diagnosis by combining global diffusion, demographic measures, and clinical ratings. Finally, to facilitate comparison with prior regional difference reports, we conducted supplementary whole-brain contrasts<sup>26</sup> of ADHD and NT participants within each sample.

## METHOD

### Participants

We report on 2 samples (1 adult and the other pediatric) obtained as part of separate studies using identical imaging protocols (Table 1). The adult sample, after quality assurance of imaging data, consisted of 42 individuals with ADHD (age range, 18.2–52.9 years, 57% male and 43% female) and 65 neurotypical (NT) comparisons (18.6–51.9 years, 65% male and 35% female). Inclusion in the adult ADHD group required a clinician's *DSM-IV-TR* diagnosis of ADHD based on the Adult ADHD Clinical Diagnostic Scale Version 1.2<sup>27</sup> and the Structured Clinical Interview for *DSM-IV*, Research Version, Non-patient Edition (SCID)<sup>28</sup> to assess Axis I disorders. Most participants with ADHD (38 of 42) met criteria for persistent ADHD diagnosis (i.e., symptoms and impairment in childhood and adulthood), 2 participants for current ADHD only (i.e., meeting criteria only in adulthood), and 2 presented with history of ADHD in remission (i.e., symptoms in only childhood). All but 7 participants completed the self-report Conners' Adult ADHD Rating Scales (CAARS).<sup>29</sup>

The pediatric sample, after imaging data quality assurance, consisted of 82 individuals with ADHD (age range, 5.2–17.2 years, 78% male and 22% female) and 80 NT children (4.9–17.7 years, 69% male and 31% female). Inclusion in the ADHD group required a clinician's *DSM-IV-TR* diagnosis of ADHD supported by review of prior history and results of the Conners' Parent Rating Scale–Revised: Long Version (CPRS-R:LV, obtained for all but 22 participants)<sup>30</sup> and psychiatric interview using the Schedule for Affective Disorders and Schizophrenia for School-Age Children–Present and Lifetime Version (KSADS-PL),<sup>31</sup> administered separately to child and parent.

Inclusion as NT in both the adult and pediatric samples required absence of current Axis I diagnosis, assessed for the adult sample with SCID, and for pediatric sample with KSADS-PL (administered to both child and parent in 75 instances; in 5 instances, child-only KSADS-PL was supplemented by unstructured clinical interviews of the parent). Exclusion criteria for all participants were current evidence of autism, major depression, suicidality, substance-related disorder, obsessive-compulsive disorder, conduct disorder, post-traumatic stress disorder, panic disorder, Tourette disorder, lifetime history of psychosis or mania, general chronic medical conditions, left-handedness, or estimated full-scale IQ below 80. Comorbid disorders were present in 7 adults and 27 children with ADHD (see Supplement 1, available online). The Wechsler Abbreviated Scale of Intelligence (WASI) provided estimates of full-scale IQ<sup>32</sup> in all adults and all but 20 children (Differential Ability Scales II<sup>33</sup> was used for 1 NT child, and the Kaufman Brief Intelligence Test<sup>34</sup> for 3 NT and 16

children with ADHD). The institutional review boards of the New York University (NYU) School of Medicine and NYU granted ethical approval. All participants provided written informed consent and, for minors, assent.

### Data Acquisition

Magnetic resonance imaging (MRI) data were obtained at the NYU Center for Brain Imaging using a 3T Siemens Allegra scanner with a single-channel Nova head coil. Anatomical T1-weighted images were obtained using 3D Magnetization Prepared Rapid Acquisition Gradient Echo sequence (TR = 2,530 milliseconds; TE = 3.25 milliseconds; TI = 1,100 milliseconds; flip angle = 7°). Two DTI scans were acquired using a twice-refocused diffusion-weighted echo-planar image sequence with parameters TR = 5,200 milliseconds; TE = 78 milliseconds; 50 slices; acquisition matrix 64 × 64; field of view = 92 mm; acquisition voxel size = 3 × 3 × 3 mm; 64 non-collinear diffusion directions, uniformly distributed around a unit sphere with a b-value of 1000 s/mm<sup>2</sup>; 1 image with no diffusion weighting. A gradient echo field map was collected (TR = 834 milliseconds; TEs = 5.23 and 7.69 milliseconds) with slice position and resolution identical to those of the diffusion-weighted images.

### DTI Preprocessing and Quality Assurance

Diffusion-weighted data were pre-processed using FMRIB Software Library version 5.<sup>35</sup> Motion correction (linear registration) was followed by correction of image distortions from eddy currents and B0-field inhomogeneities. Absolute intervolume displacement<sup>36</sup> of each image with respect to the first image in the run was computed. For participants with maximum displacement within 1.5 × voxel size (grand total = 410 participants), individual maps were visually inspected for signal dropout, brain coverage, artifacts, and additional motion. Based on quality control criteria, data from 5 ADHD (11% of initial sample of adults with ADHD) and 12 NT (13%) adults, and 67 ADHD (45%) and 57 NT (42%) children, were discarded. Mean absolute intervolume displacement served as the primary head motion index for subsequent analyses. Analyses were repeated with mean volume-by-volume translation and mean volume-by-volume rotation as supplementary motion indices.<sup>36,37</sup> Diffusion gradients were rotated to improve consistency with the motion parameters, and data for each of the two 64-direction scans were used to fit the tensor parameters, thus improving signal-to-noise ratio. Following fitting of diffusion tensors at each voxel, fractional anisotropy (FA), axial diffusivity (AD), radial diffusivity (RD), mean diffusivity (MD), and mode of anisotropy (MA) values were generated. Three additional shape measures (linear  $c_L$ , planar  $c_P$ , and spherical  $c_S$  tensor components<sup>38</sup>) were computed to characterize post hoc differences in diffusion tensor geometry.<sup>19</sup> All analyses detailed below (global and voxelwise) were performed separately for the pediatric and adult samples, unless noted otherwise.

### Tract-Based Spatial Statistics (TBSS)

Further data processing was performed using tract-based spatial statistics (TBSS),<sup>25</sup> which calculates a white matter “skeleton” to represent the center of each white-matter tract common to all participants, thus ameliorating the impact of imperfect alignment, registration, and arbitrarily thresholded spatial smoothing. TBSS is commonly preferred

when contrasting patients with NT individuals despite potential drawbacks.<sup>39</sup> First, nonlinear registration aligned every FA image to every other one, identifying the “most representative” image as the target. This target was then affine-aligned into Montreal Neurological Institute (MNI) 152 standard space, and each image transformed into  $1 \times 1 \times 1\text{-mm}^3$  MNI space by combining the nonlinear transform to the target with the affine transform from that target to MNI152 space. Next, each participant’s images were brought into the standard space (study-specific template, separate for each of the 2 samples) via nonlinear transformations. Voxelwise and global analyses were conducted for voxels within the extent of the mean FA skeleton (FA threshold = 0.3) with FA, AD, RD, MD, and MA projected onto it.

## Analyses

**Group Characteristics**—In each sample, the NT and ADHD groups were compared with respect to demographic, clinical, and head motion measures using  $\chi^2$  and unpaired 2-tailed  $t$  tests. As motion can elicit spurious between-group differences,<sup>36</sup> we also examined whether global MA and motion were correlated.

**Global White Matter Analysis of Diffusion Characteristics**—First, we checked for outliers. The upper and lower thresholds were set at  $2.2 \times$  interquartile range (IQR) above the 75th percentile and below the 25th percentile, respectively.<sup>40</sup> In adults, no outliers were found in FA, AD, RD, or MA. Two NT participants had MD values of  $3.1 \times$  IQR and  $2.5 \times$  IQR below the 25th percentile. Among children, one outlier was found for AD (NT at  $2.3 \times$  IQR above the 75th percentile). Outliers were excluded from subsequent global diffusion analyses involving MD or AD.

**ADHD Versus NT Contrasts**—A global index, that is, values averaged across the entire skeleton, was calculated separately for FA, AD, RD, MD, and MA, yielding 1 value per DTI index per individual. Analogously, global  $c_L$ ,  $c_P$ , and  $c_S$  were calculated. Global indices were contrasted between the ADHD and NT groups using 2-tailed  $t$  tests and linear models controlling for the common confounders age, sex, ADHD subtype, IQ, medication history (stimulant-naive or not), and head motion; effect sizes are given as Cohen’s  $d$ .

**ADHD Diagnosis Correlates**—We investigated the effect of microstructural DTI metrics and of demographic and clinical rating measures on the probability of an ADHD diagnosis. To allow direct comparison between measures,<sup>41</sup> we used a forward stepwise binary logistic regression including the global DTI indices (FA, AD, RD, MD, and MA), age, sex, motion, and the respective Conners’ ADHD total  $T$  score (adult sample: CAARS; pediatric sample: CPRS-R:LV) with default cut-off value of 0.5 and criterion of adding ( $p < .05$ ) and keeping only significant correlates ( $p < .1$ ). All global diffusion indices were linearly scaled by a factor of 100 to facilitate interpretation of unit change in odds ratio. Secondly, effects of predictors were illustrated using receiver operating characteristic (ROC) curves and 10-fold cross-validated area under the ROC curves.

**Confirmatory Analyses in Native Space**—Since this is the first study to examine global MA effects in ADHD, supplementary analyses were carried out to examine the

possibility that TBSS registration and normalization procedures<sup>39</sup> might affect MA and inflate effect sizes. Accordingly, to evaluate MA effects independently of the biases inherent to registration to a common template, global MA for each participant was also computed within each individual's native space. We derived white matter native space boundaries using FA, rather than MA, to provide an independent estimate from a mathematically orthogonal measure. After assessing 3 potential FA thresholds in the 0.40–0.45 range, individual FA maps were thresholded at  $>0.43$ , a cut-off value that concurrently satisfied 2 conditions: predominant inclusion of white matter, and preservation of a minimal volume comparable in size to the one derived from TBSS. We first confirmed that the volumes of the resulting individual masks did not differ between the ADHD and NT groups in either sample ( $p > .2$ ). All analyses conducted with TBSS-derived MA were repeated using MA computed in native space as defined by each individual's FA mask. Finally, we also examined the correlation between MA and age across the pediatric and adult samples combined.

**Supplementary Voxelwise Analyses**—Supplementary analyses assessed distributed, voxelwise differences across the white matter skeleton. We examined the effect of diagnosis (ADHD, NT) in separate models for FA (the most widely used diffusion index) and MA (the only DTI measure that differed significantly in both samples), adjusting for age, sex, motion, and its respective global diffusion index (given that global diffusion differed significantly between groups). To control for multiple comparisons (familywise error rate at  $\alpha = 0.05$ ), we applied the threshold-free cluster enhancement algorithm,<sup>26</sup> with standard settings for DTI.

## RESULTS

### Group Characteristics

Demographic factors (e.g., IQ) were equivalent between the ADHD and NT groups for both samples (Table 1). Head motion did not differ significantly between NT participants and those with ADHD in either sample ( $p_{\min} = 0.21$ ), with mean rotation and translation values comparable to those in previous reports.<sup>36,37</sup> In addition, motion was not significantly correlated with global MA in either diagnostic group in either of the samples ( $p_{\min}=0.45$ ; see Figure S2, available online).

### Global White Matter Analysis of Microstructural Characteristics

We first directly contrasted ADHD and NT in each age sample separately for every global DTI metric. As depicted in Figure 1A, global FA, AD, MD, and MA were significantly lower for the adult group with ADHD relative to the corresponding NT group, with greatest effect size for MA ( $d = 1.45$ ). For the pediatric sample, only MA was significantly lower in ADHD ( $d = 0.68$ ) (Figure 1B). Our post hoc analyses of linear, planar, and spherical shape measures largely supported the effects observed using MA. Planar anisotropy was significantly greater in individuals with ADHD relative to NT in both samples. Adults with ADHD also had greater spherical anisotropy and lower linear anisotropy than NT adults (Table S1, available online). Figure 2 shows the distributions of global MA values for both samples.

Global MA differences in each of the 2 samples suggest that MA is associated with ADHD diagnosis. Nonetheless, we considered possible confounds. First, we repeated analyses with age, sex, head motion, ADHD subtype, IQ, and medication history as covariates. The pattern of results and effect sizes were essentially unchanged (data not shown). Second, to address possible biases from the transformations inherent to spatial normalization via TBSS, we repeated analyses in native space using white matter masks based on each individual's FA. Lower MA in ADHD in native space was corroborated for adult ( $t_{105} = 7.45, p < .001, d = 1.40$ ) and pediatric samples ( $t_{160} = 5.13, p < .001, d = 0.62$ ). We also examined the subset of 38 adult participants reporting persistent ADHD and confirmed significantly lower global MA relative to NT ( $t_{101} = 6.70, p < .001, d = 1.37$ ).

By design, our adult and pediatric samples were analyzed separately. Still, as the 2 samples had nonoverlapping age ranges, we explored the relationship between global MA and age by aggregating data across samples (see Figure S3, available online). A significant effect of age was found for the primary global MA measure derived from the TBSS skeleton ( $t_{267} = 3.32, p < .05$ ) and the confirmatory global MA measure derived in native space ( $t_{267} = 2.04, p < .05$ ). As expected, adjusting for age, the main effect of diagnosis was corroborated in the combined sample ( $p_{\max} = .005$ ) with no significant interaction between age and diagnosis ( $p_{\min} = .16$ ). Further analyses with penalized splines<sup>42</sup> did not detect nonlinear age effects or age-varying ADHD effects (data not shown).

The relation between global MA and Conners' scores for each sample is illustrated in Figure S4 (available online). Examined within each group separately, 2-tailed Pearson correlations between global MA and Conners' scores did not reach significance ( $p_{\min} = .22$ ). Across all individuals within a sample, Conners' scores were significantly correlated with global MA (adult sample:  $r_{98} = -0.505, p < .001$ ; pediatric sample:  $r_{138} = -0.274, p < .001$ ). We found no evidence of global MA differences between the 2 most predominant ADHD subtypes in our samples, Combined type and Inattentive type (adult sample:  $t_{39} = 1.58, p = .12$ , pediatric sample:  $t_{74} = 0.62, p = .54$ ). Finally, we corroborated that global MA values were comparable between medication-naïve patients and those with a history of stimulant use (adult sample:  $t_{40} = 0.48, p = .63$ , pediatric sample:  $t_{80} = 0.68, p = .50$ ).

### Predictors of ADHD Diagnosis

In line with the long-term goal of developing biomarkers,<sup>43</sup> we carried out forward stepwise binary logistic regression (Table 2), which included all global DTI measures, ADHD rating scales, age, sex, and motion, to identify the maximally predictive combination of variables for ADHD separately within each sample. Stepwise logistic regression yielded a model that contained only Conners' scales (CAARS or CPRS-R:LV, for adult and pediatric samples, respectively) and global MA as significant predictors. In the pediatric sample, analyses using Conners' Teachers Rating Scale-Revised: Long Version<sup>30</sup> scores corroborated results obtained with CPRS-R:LV, as shown in Supplement 1, available online. The ROC curves for Conners' scores, with and without global MA in the model, are depicted in Figure S5, available online. Sensitivity and specificity and positive and negative predictive values for the significant predictors of ADHD diagnosis in each step are reported in Table S2, available online.

In addition, in each sample, we plotted ROC curves for the global diffusion indices that differed between ADHD and NT individuals, used as correlates of ADHD diagnosis (Figure 3), and computed the area under the ROC curve for every global DTI index (see Table S3, available online). Consistent with its effect sizes, MA had substantially greater predictive power for diagnosis than any of the other measures in either sample (adult MA area under the ROC = 0.88, pediatric MA area under the ROC = 0.70).

### Voxelwise Whole-Brain Analyses

No main effect of diagnosis (ADHD > NT or ADHD < NT) was found for FA or MA when controlling type I error for whole-brain multiple comparisons (threshold-free cluster enhancement–corrected voxelwise, all yielding  $p > .1$ ) for adults or children. The FA null effect is consistent with a recent report.<sup>44</sup>

## DISCUSSION

We compared global, whole-brain differences in diffusion-based measures of white matter between individuals with and without ADHD. Primary analyses were carried out in an adult sample, and corroboratory analyses in a separate pediatric sample. In both samples, we found the following: (1) among global diffusion metrics, the mode of anisotropy (MA)<sup>19</sup> exhibited the greatest effect size for differences between ADHD and comparison participants (with lower values for ADHD than NT); and (2) global MA significantly improved predictive accuracy for ADHD diagnosis, above the substantial predictive utility of clinical measures. In the adult sample, global FA, AD, and MD were also significantly lower in patients but with smaller effect sizes. We did not observe significant regional between-group differences in FA in either sample after correction for multiple comparisons.

Synthesis of the mosaic of regional findings accumulating from ADHD diffusion studies remains challenging.<sup>4</sup> Despite our null voxelwise results, we observed strong global effects. Global differences in diffusion properties between ADHD and NT groups might arise in several ways. First, differential diagnostic effects might inhere in large-scale distributed networks,<sup>1</sup> which we were unable to detect because of our limited voxel resolution and sample sizes. Second, different circuits might be affected to a greater extent in some individuals than in others, resulting in an apparently global effect at the group level. Third, small but consistent changes in microstructure may be uniformly distributed throughout white matter in ADHD, possibly coexisting with regional alterations. Such nuanced interpretations fit with the widespread deficiency of ADHD connectivity proposed by recent models.<sup>5</sup> We did not find significant global white or gray matter volumetric between-group differences in either sample (data not shown). One may speculate that, given the positive correlation between total brain volume and IQ<sup>45</sup> and the fact that our diagnostic groups did not significantly differ in IQ, this is not surprising. The lack of global white matter volumetric effects is consistent with a recent large-scale study<sup>7</sup> and broadly suggests that MA effects are not reducible to gross morphometric alterations between groups.

The strongest global diffusion effect that we observed was reduced MA in ADHD, which indicates greater planar anisotropy at the expense of linear anisotropy. This suggests a preponderance of more complex white matter architecture, for example, crossing fibers,



within an average voxel, consistent with a greater disorganization of white matter fibers<sup>24,46</sup> in ADHD. Empirically relating changes in MA to underlying white matter properties is in its infancy. However, robust MA differences across various ages were detected in an animal model of a neurodegenerative disease<sup>46</sup>; MA was useful in further characterizing differences in FA<sup>21</sup>; and MA significantly improved the ability to prospectively predict which individuals would develop Alzheimer's disease.<sup>24</sup>

Although the effect size and predictive benefit of MA was greater in adults, the confidence intervals overlapped across samples. Furthermore, we did not observe a significant age-by-diagnosis interaction between MA and age, and MA was only modestly, albeit significantly, related to age ( $R^2 = 0.04$ ).

Our results should be considered in light of limitations. First, replication using independently collected data with smaller voxel sizes is needed to establish generalizability. Second, residual head motion effects cannot be ruled out; although we corrected for head motion after data quality assurance, our motion estimates fell within recently published ranges,<sup>36</sup> and motion did not differ between diagnostic groups or correlate with MA. Third, we confirmed the principal between-group effects for both samples using additional tensor component indices to characterize ellipsoid shape<sup>38</sup>; still, MA estimates in areas with low signal are vulnerable to artifacts and image distortion.<sup>19</sup> Fourth, we included patients who were medication-naïve as well as recently medicated, although we found no evidence that this affected our results. Fifth, unlike most studies in ADHD, our groups did not differ significantly in IQ, which may have contributed to the lack of between-group volumetric differences. Finally, although we observed ADHD-related differences in MA with effect sizes approaching the levels necessary for diagnostic biomarker candidates, at least in adults,<sup>43</sup> we cannot address their clinical specificity. Future studies should compare global MA between ADHD and other disorders. Fortunately, such analyses can be readily performed with already-collected DTI data, as computing MA in standard software is straightforward.

Following the technological and methodological innovations of the Human Connectome Project<sup>47</sup> (e.g., multiband imaging and increased gradient strength, which improve sampling of the diffusion space and spatial resolution), diffusion imaging methods are advancing. Superior data acquisition will enhance the accuracy and precision of DTI metrics and thus will offer better sensitivity in differentiating between individuals and diagnostic groups. In the meantime, our results underscore the value of applying novel analytic approaches to existing data (especially larger aggregated datasets).<sup>48</sup> Specifically, we highlight the power of 2 relatively uncommon strategies—use of global measures and of complementary diffusion metrics, including MA—to augment conventional DTI analyses and to illuminate pathophysiology.

## Supplementary Material

Refer to Web version on PubMed Central for supplementary material.

## Acknowledgments

This work was supported by R01MH081218, R01MH083246, R01HD065282, R01MH091140, T32MH067763, and grants from the Stavros Niarchos Foundation.

The authors gratefully acknowledge all the participants and their parents, Pablo Velasco, PhD, and Keith Sanzenbach, R, RT, MR, of the NYU Center for Brain Imaging, for support in developing imaging sequences, Stephen Smith, PhD, and Gwenaëlle Douaud, PhD, of Oxford University, for helpful discussion of the mode of anisotropy, and Amy K. Roy, PhD, of Fordham University, for providing data from an ongoing study of children with ADHD and comparisons (R01MH091140).

## References

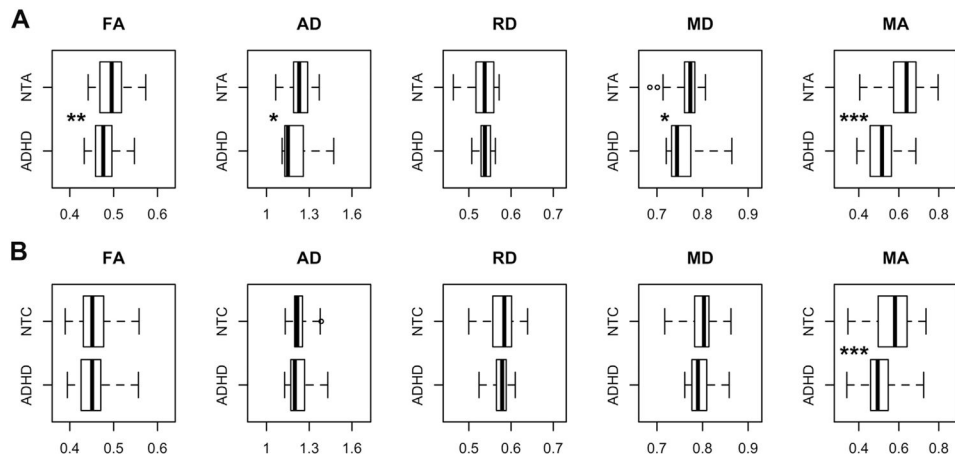
1. Castellanos FX, Proal E. Large-scale brain systems in ADHD: beyond the prefrontal-striatal model. *Trends Cogn Sci*. 2012; 16:17–26. [PubMed: 22169776]
2. Liston C, Cohen MM, Teslovich T, Levenson D, Casey BJ. Atypical pre-frontal connectivity in attention-deficit/hyperactivity disorder: pathway to disease or pathological end point? *Biol Psychiatry*. 2011; 69:1168–1177. [PubMed: 21546000]
3. Konrad K, Eickhoff SB. Is the ADHD brain wired differently? A review on structural and functional connectivity in attention deficit hyperactivity disorder. *Hum Brain Mapp*. 2010; 31:904–916. [PubMed: 20496381]
4. van Ewijk H, Heslenfeld DJ, Zwiers MP, Buitelaar JK, Oosterlaan J. Diffusion tensor imaging in attention deficit/hyperactivity disorder: a systematic review and meta-analysis. *Neurosci Biobehav Rev*. 2012; 36:1093–1106. [PubMed: 22305957]
5. Castellanos FX, Proal E. Location, location, and thickness: volumetric neuroimaging of attention-deficit/hyperactivity disorder comes of age. *J Am Acad Child Adolesc Psychiatry*. 2009; 48:979–981. [PubMed: 20854767]
6. Krain AL, Castellanos FX. Brain development and ADHD. *Clin Psychol Rev*. 2006; 26:433–444. [PubMed: 16480802]
7. Greven CU, Bralten J, Mennes M, et al. Developmentally stable whole-brain volume reductions and developmentally sensitive caudate and putamen volume alterations in those with attention-deficit/hyperactivity disorder and their unaffected siblings. *JAMA Psychiatry*. 2015; 72:490–499. [PubMed: 25785435]
8. Makris N, Biederman J, Valera EM, et al. Cortical thinning of the attention and executive function networks in adults with attention-deficit/hyperactivity disorder. *Cereb Cortex*. 2007; 17:1364–1375. [PubMed: 16920883]
9. Proal E, Reiss PT, Klein RG, et al. Brain gray matter deficits at 33-year follow-up in adults with attention-deficit/hyperactivity disorder established in childhood. *Arch Gen Psychiatry*. 2011; 68:1122–1134. [PubMed: 22065528]
10. Shaw P, Lerch J, Greenstein D, et al. Longitudinal mapping of cortical thickness and clinical outcome in children and adolescents with attention-deficit/hyperactivity disorder. *Arch Gen Psychiatry*. 2006; 63:540–549. [PubMed: 16651511]
11. Francx, W.; Zwiers, MP.; Mennes, M., et al. White matter microstructure and developmental improvement of hyperactive/impulsive symptoms in attention-deficit/hyperactivity disorder. *J Child Psychol Psychiatry*. 2015 Jan 10. <http://dx.doi.org/10.1111/jcpp.12379> [Epub ahead of print]
12. van Ewijk H, Heslenfeld DJ, Zwiers MP, et al. Different mechanisms of white matter abnormalities in attention-deficit/hyperactivity disorder: a diffusion tensor imaging study. *J Am Acad Child Adolesc Psychiatry*. 2014; 53:790–799.e3. [PubMed: 24954828]
13. Onnink AMH, Zwiers MP, Hoogman M, et al. Deviant white matter structure in adults with attention-deficit/hyperactivity disorder points to aberrant myelination and affects neuropsychological performance. *Prog Neuropsychopharmacol Biol Psychiatry*. 2015; 63:14–22. [PubMed: 25956761]
14. Casey BJ, Epstein JN, Buhle J, et al. Frontostriatal connectivity and its role in cognitive control in parent-child dyads with ADHD. *Am J Psychiatry*. 2007; 164:1729–1736. [PubMed: 17974939]

15. Reiss PT, Schwartzman A, Lu F, Huang L, Proal E. Paradoxical results of adaptive false discovery rate procedures in neuroimaging studies. *Neuroimage*. 2012; 63:1833–1840. [PubMed: 22842214]
16. Basser PJ, Pierpaoli C. Microstructural and physiological features of tissues elucidated by quantitative-diffusion-tensor MRI. *J Magn Reson B*. 1996; 111:209–219. [PubMed: 8661285]
17. Jones DK, Knösche TR, Turner R. White matter integrity, fiber count, and other fallacies: the do's and don'ts of diffusion MRI. *Neuroimage*. 2013; 73:239–254. [PubMed: 22846632]
18. Jeurissen B, Leemans A, Tournier J-D, Jones DK, Sijbers J. Investigating the prevalence of complex fiber configurations in white matter tissue with diffusion magnetic resonance imaging. *Hum Brain Mapp*. 2013; 34:2747–2766. [PubMed: 22611035]
19. Ennis DB, Kindlmann G. Orthogonal tensor invariants and the analysis of diffusion tensor magnetic resonance images. *Magn Reson Med*. 2006; 55:136–146. [PubMed: 16342267]
20. Kindlmann G, Ennis DB, Whitaker RT, Westin CF. Diffusion tensor analysis with invariant gradients and rotation tangents. *IEEE Trans Med Imaging*. 2007; 26:1483–1499. [PubMed: 18041264]
21. Douaud G, Jbabdi S, Behrens TEJ, et al. DTI measures in crossing-fibre areas: increased diffusion anisotropy reveals early white matter alteration in MCI and mild Alzheimer's disease. *Neuroimage*. 2011; 55:880–890. [PubMed: 21182970]
22. Itahashi T, Yamada T, Nakamura M, et al. Linked alterations in gray and white matter morphology in adults with high-functioning autism spectrum disorder: a multimodal brain imaging study. *NeuroImage Clin*. 2014; 7:155–169. [PubMed: 25610777]
23. Wigand M, Kubicki M, Clemm von Hohenberg C, et al. Auditory verbal hallucinations and the interhemispheric auditory pathway in chronic schizophrenia. *World J Biol Psychiatry*. 2015; 16:31–44. [PubMed: 25224883]
24. Douaud G, Menke RAL, Gass A, et al. Brain microstructure reveals early abnormalities more than two years prior to clinical progression from mild cognitive impairment to Alzheimer's disease. *J Neurosci*. 2013; 33:2147–2155. [PubMed: 23365250]
25. Smith SM, Jenkinson M, Johansen-Berg H, et al. Tract-based spatial statistics: voxelwise analysis of multi-subject diffusion data. *Neuroimage*. 2006; 31:1487–1505. [PubMed: 16624579]
26. Smith SM, Nichols TE. Threshold-free cluster enhancement: addressing problems of smoothing, threshold dependence and localisation in cluster inference. *NeuroImage*. 2009; 44:83–98. [PubMed: 18501637]
27. Adler, L.; Spencer, T. *The Adult ADHD Clinical Diagnostic Scales (ACDS)*. New York: New York University, School of Medicine; 2004.
28. First, M.; Spitzer, R.; Gibbon, M.; Williams, J. *Structured Clinical Interview for DSM-IV-TR Axis I Disorders, Research Version, Non-Patient Edition*. New York: Biometrics Research, New York State Psychiatric Institute; 2002. SCID-I/NP
29. Conners, C.; Erhardt, D.; Sparrow, E. *Conners' Adult ADHD Rating Scales: Technical Manual*. Toronto: Multi-Health Systems; 1999.
30. Conners, C. *Conners' Rating Scales—Revised User's Manual*. Tonawanda, NY: Multi-Health Systems; 1997.
31. Kaufman J, Birmaher B, Brent D, et al. Schedule for Affective Disorders and Schizophrenia for School-Age Children—Present and Lifetime Version (K-SADS-PL): initial reliability and validity data. *J Am Acad Child Adolesc Psychiatry*. 1997; 36:980–988. [PubMed: 9204677]
32. Wechsler, D. *Wechsler Abbreviated Scale of Intelligence (WASI)*. San Antonio, TX: Psychological Corporation; 1999.
33. Elliott, C. *Differential Ability Scales®-II*. San Antonio, TX: Harcourt Assessment; 2007.
34. Kaufman, A.; Kaufman, N. *Kaufman Brief Intelligence Test*. Circle Pines, MN: American Guidance Service; 1990.
35. Smith SM, Jenkinson M, Woolrich MW, et al. Advances in functional and structural MR image analysis and implementation as FSL. *Neuroimage*. 2004; 23(Suppl 1):S208–S219. [PubMed: 15501092]
36. Yendiki A, Koldewyn K, Kakunoori S, Kanwisher N, Fischl B. Spurious group differences due to head motion in a diffusion MRI study. *Neuroimage*. 2014; 88:79–90.

37. Hong S-B, Zalesky A, Fornito A, et al. Connectomic disturbances in attention-deficit/hyperactivity disorder: a whole-brain tractography analysis. *Biol Psychiatry*. 2014; 76:656–663. [PubMed: 24503470]
38. Westin CF, Maier SE, Mamata H, Nabavi A, Jolesz FA, Kikinis R. Processing and visualization for diffusion tensor MRI. *Med Image Anal*. 2002; 6:93–108. [PubMed: 12044998]
39. Bach M, Laun FB, Leemans A, et al. Methodological considerations on tract-based spatial statistics (TBSS). *Neuroimage*. 2014; 100:358–369. [PubMed: 24945661]
40. Hoaglin DC, Iglewicz B. Fine-tuning some resistant rules for outlier labeling. *J Am Stat Assoc*. 1987; 82:1147–1149.
41. Pepe MS, Kerr KF, Longton G, Wang Z. Testing for improvement in prediction model performance. *Stat Med*. 2013; 32:1467–1482. [PubMed: 23296397]
42. Wood, S. *Generalized Additive Models: An Introduction with R*. Boca Raton, FL: Chapman and Hall; 2006.
43. Castellanos FX, Di Martino A, Craddock RC, Mehta AD, Milham MP. Clinical applications of the functional connectome. *Neuroimage*. 2013; 80:527–540. [PubMed: 23631991]
44. Cooper M, Thapar A, Jones DK. White matter microstructure predicts autistic traits in attention-deficit/hyperactivity disorder. *J Autism Dev Disord*. 2014; 44:2742–2754. [PubMed: 24827086]
45. Lefebvre A, Beggiato A, Bourgeron T, Toro R. Neuroanatomical diversity of corpus callosum and brain volume in autism: meta-analysis, analysis of the autism brain imaging data exchange project, and simulation. *Biol Psychiatry*. 2015; 78:126–134. [PubMed: 25850620]
46. Sahara N, Perez PD, Lin W-L, et al. Age-related decline in white matter integrity in a mouse model of tauopathy: an in vivo diffusion tensor magnetic resonance imaging study. *Neurobiol Aging*. 2014; 35:1364–1374. [PubMed: 24411290]
47. Van Essen DC, Ugurbil K, Auerbach E, et al. The Human Connectome Project: a data acquisition perspective. *Neuroimage*. 2012; 62:2222–2231. [PubMed: 22366334]
48. Thompson PM, Stein JL, Medland SE, et al. The ENIGMA Consortium: large-scale collaborative analyses of neuroimaging and genetic data. *Brain Imaging Behav*. 2014; 8:153–182. [PubMed: 24399358]

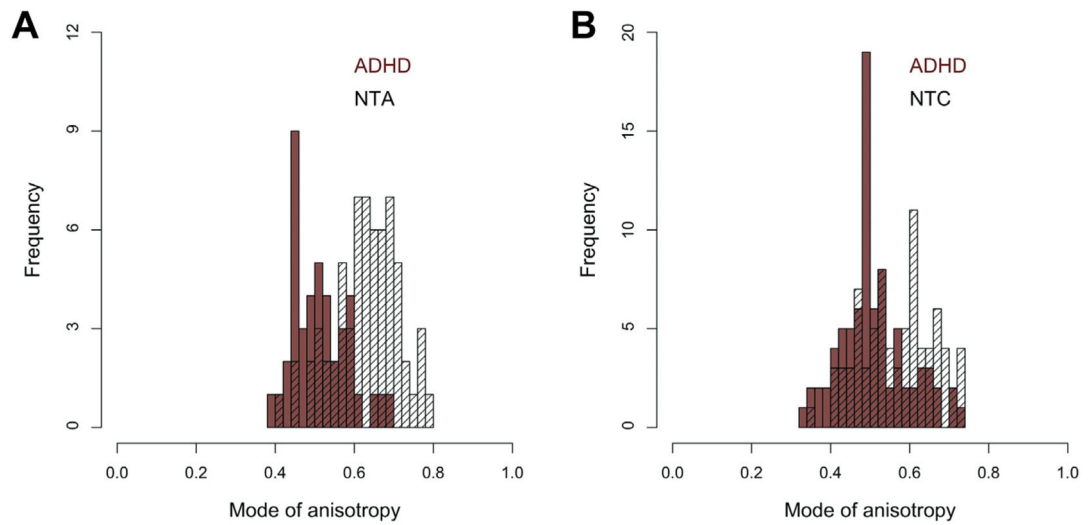
### Clinical Guidance

- Global diffusion measures, particularly the mode of anisotropy, differed between individuals with and without ADHD in two separate samples, which suggests a global white matter alteration in ADHD.
- The effect size of global mode of anisotropy reduction in ADHD approaches the acceptable range for candidate biomarkers.
- Despite these preliminary results, diffusion indices do not provide specific clinical guidance regarding individual patients and are not yet recommended as part of standard clinical care.

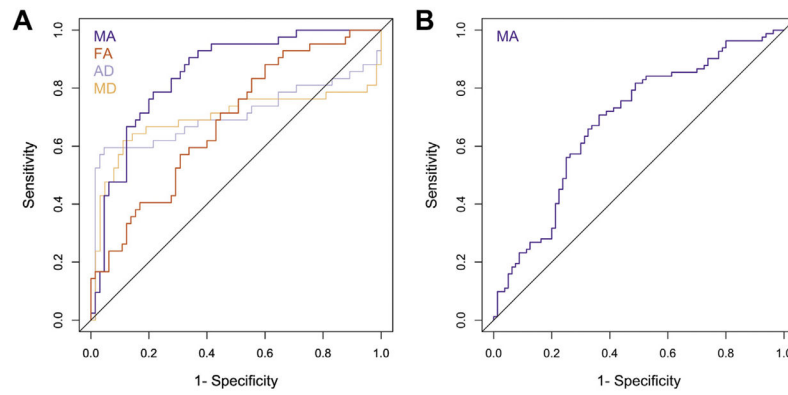


**FIGURE 1.**

Box plots for global diffusion measures in (A) adult sample and (B) pediatric sample. Note: Significant differences between individuals with attention-deficit/hyperactivity disorder (ADHD) and neurotypical adults (NTA) or neurotypical children (NTC) are denoted with \* $p < .05$ , \*\* $p < .01$ , \*\*\* $p < .001$ . Outliers ( $n = 3$ ) are plotted but excluded from  $t$  tests. See Table S1, available online, for summary statistics and effect sizes. AD = axial diffusivity; FA = fractional anisotropy; MA = mode of anisotropy; MD = mean diffusivity; RD = radial diffusivity. Units for AD, RD, and MD are  $10^{-3}\text{mm}^2/\text{s}$ .

**FIGURE 2.**

Histograms of average global mode of anisotropy values in individuals from (A) adult sample and (B) pediatric sample. Note: ADHD = attention-deficit/hyperactivity disorder; NTA = neurotypical adults; NTC = neurotypical children.



**FIGURE 3.**

Receiver operating characteristic (ROC) curves for the global diffusion indices that differed between individuals with attention-deficit/hyperactivity disorder (ADHD) and neurotypical (NT) individuals in (A) adult sample and (B) pediatric sample. Note: The ROC curve plots the true positive rate versus the false-positive rate as the binary classification threshold varies. Classification was conducted separately for each diffusion measure. AD = axial diffusivity; FA = fractional anisotropy; MA = mode of anisotropy; MD = mean diffusivity.



**TABLE 1**  
Demographic and Clinical Characteristics of Participants With Analyzable Diffusion Tensor Imaging (DTI) Data

	Adult Sample			Pediatric Sample		
	ADHD	NT	ADHD vs. NT	ADHD	NT	ADHD vs. NT
Sex, n (male, female)	24, 18	42, 23	$\chi^2 = 0.61, p = .44$	64, 18	55, 25	$\chi^2 = 1.80, p = .18$
Age (y)	31.65 ± 9.8	31.06 ± 9.0	$t_{105} = 0.32, p = .75$	10.63 ± 2.8	11.04 ± 2.6	$t_{160} = -0.96, p = .34$
Full IQ	111.1 ± 11.8	110.9 ± 10.6	$t_{105} = 0.11, p = .92$	105.52 ± 14.9	108.98 ± 14.3	$t_{160} = -1.51, p = .13$
Race/ethnicity <sup>a</sup>						
Caucasian	23 (59)	30 (48)	$\chi^2 = 5.52, p = .06$	43 (57)	40 (54)	$\chi^2 = 0.23, p = .89$
African-American	4 (10)	19 (30)		17 (22)	19 (26)	
Other <sup>b</sup>	12 (31)	14 (22)		16 (21)	15 (20)	
Medication status: n (%)						
Medication naïve	14 (33)	—	—	59 (72)	—	—
Not naïve, off medication(s) <sup>c</sup>	26 (62)	—	—	18 (22)	—	—
Current stimulant medication <sup>d</sup>	2 (5)	—	—	4 (5)	—	—
Current nonstimulant medication <sup>e</sup>	0 (0)	—	—	1 (1)	—	—
Head motion indices						
Absolute motion (mm)	2.82 ± 0.58	2.75 ± 0.58	$t_{105} = 0.58, p = .56$	1.98 ± 0.40	2.07 ± 0.41	$t_{160} = -1.38, p = .17$
Volume-by-volume translation (mm)	0.91 ± 0.25	0.87 ± 0.24	$t_{105} = 0.91, p = .37$	0.75 ± 0.28	0.73 ± 0.16	$t_{160} = 0.47, p = .64$
Volume-by-volume rotation (degrees)	0.29 ± 0.11	0.33 ± 0.18	$t_{105} = -0.96, p = .34$	0.26 ± 0.07	0.26 ± 0.09	$t_{160} = 0.14, p = .89$
CAARS (T score)						
DSM-IV Inattentive	73.11 ± 11.85	41.47 ± 9.29	$t_{98} = 14.89, p < .001$			
DSM-IV Hyperactive-Impulsive	61.16 ± 13.61	38.87 ± 8.11	$t_{53.3} = 10.28, p < .001$			
DSM-IV Total Symptoms	69.89 ± 12.24	38.66 ± 9.66	$t_{98} = 14.16, p < .001$			
CPRS-R:LV (T score)						
DSM-IV Inattentive				72.41 ± 8.79	45.51 ± 6.70	$t_{120.8} = 20.18, p < .001$
DSM-IV Hyperactive-Impulsive				68.80 ± 12.70	47.07 ± 6.46	$t_{94.0} = 12.53, p < .001$
DSM-IV Total Symptoms				72.65 ± 9.26	46.03 ± 6.37	$t_{113.5} = 19.60, p < .001$
ADHD subtype: n (%)						

	Adult Sample		Pediatric Sample	
	ADHD	NT	ADHD vs. NT	ADHD vs. NT
Combined	23 (55)		50 (61)	
Predominantly Inattentive	18 (43)		26 (32)	
Predominantly Hyperactive–Impulsive	0 (0)		4 (5)	
Not otherwise specified	1 (2)		2 (2)	

Note: Continuous variables are presented as mean ± SD. ADHD = attention-deficit/hyperactivity disorder; CAARS = Conners' Adult ADHD Rating Scales; CPRS-R:LV = Conners' Parent Rating Scale–Revised: Long Version; NT = neurotypical.

<sup>a</sup>Self (parent) reported race/ethnicity was missing in 3 adults with ADHD, 2 NT adults, 6 children with ADHD, and 6 NT children.

<sup>b</sup>“Other” includes Asian, Pacific Islander, Hawaiian, American Indian, Alaskan Native, other races, or mixed races.

<sup>c</sup>Not taking psychoactive medications for periods ranging from 1 week to 1 year before scan.

<sup>d</sup>Currently treated with psychostimulants but had discontinued medication for 24 hours before scan.

<sup>e</sup>Selective serotonin reuptake inhibitor.

Binary Logistic Regression Modeling Attention-Deficit/Hyperactivity Disorder (ADHD) Diagnosis as a Function of Conners' T Scores and Global Mode of Anisotropy in Adult Sample and Pediatric Sample

TABLE 2

Adult sample	B	SE	Wald	p	OR	95% CI for OR
Step 1						
CAARS	0.23	0.05	19.85	.000	1.262	1.139–1.398
Constant	-12.91	2.89	19.91	.000		
Step 2						
CAARS	0.29	0.09	11.21	.001	1.33	1.125–1.571
Mode of anisotropy	-0.36	0.14	6.49	.011	0.7 <sup>a</sup>	0.532–0.921
Constant	4.54	5.60	0.66	.418		
<b>Pediatric sample</b>						
Step 1						
CPRS-R:LV	0.34	0.07	24.21	.000	1.407	1.128–1.613
Constant	-19.90	4.06	24.03	.000		
Step 2						
CPRS-R:LV	0.41	0.10	16.61	.000	1.504	1.236–1.831
Mode of anisotropy	-0.17	0.07	6.12	.013	0.842 <sup>b</sup>	0.734–0.965
Constant	-14.69	4.63	10.08	.002		

Note: CAARS = Conners' Adult ADHD Rating Scales; CPRS-R:LV = Conners' Parent Rating Scale-Revised; Long Version; OR = odds ratio; SE = standard error.

<sup>a</sup>The odds of ADHD diagnosis in the adult sample increase by a factor of 1.43 for each 0.01 unit reduction in the mode of anisotropy.

<sup>b</sup>The odds of ADHD diagnosis in the pediatric sample increase by a factor of 1.19 for each 0.01 unit reduction in the mode of anisotropy.